6.1 Occupancy Problem

Bins and Balls Throw n balls into n bins at random.

- 1. $\Pr[\text{Bin 1 is empty}] = (1 \frac{1}{n})^n \sim \frac{1}{e}.$
- 2. $\Pr[\text{Bin 1 has k balls}] = \binom{n}{k} \frac{1}{n}^k (1 \frac{1}{n})^{n-k} \le \frac{1}{e \cdot k!}.$

Sterling's Approximations

$$(\frac{n}{k})^k \le \binom{n}{k} \le (\frac{ne}{k})^k$$

Thus, letting $A_{i,k}$ be the event that bin *i* contains at least *k* balls, we have

$$\mathbf{Pr}(A_{i,k}) = \sum_{i=k}^{n} {\binom{n}{i}^{i} \left(\frac{i}{n}\right)^{i} \left(1-\frac{i}{n}\right)^{n-k}}$$

Thus, by the union bound,

$$\mathbf{Pr}(\text{any bin contains more than } k \text{ balls}) \leq \sum_{i=1}^{n} \mathbf{Pr}(A_{i,k})$$

In order to approximate this, we need to derive a simple upper bound for $\mathbf{Pr}(A_{i,k})$. We'll make use of the following elementary inequality, for any $i \leq n$:

$$\left(\frac{n}{i}\right)^i \le \binom{n}{i} \le \left(\frac{ne}{i}\right)^i$$

Using this we can easily derive the bound

$$\mathbf{Pr}(A_{i,k}) \leq \sum_{i=k}^{n} \left(\frac{ne}{i}\right)^{i} \left(\frac{1}{n}\right)^{i}$$
$$= \left(\frac{e}{i}\right)^{k} \left(1 + \frac{e}{k} + \left(\frac{e}{k}\right)^{2} + \cdots\right)$$
$$= \left(\frac{e}{k}\right)^{k} \frac{1}{1 - e/k}$$

Now comes the tedious part. Let $k = \lceil (3 \log n) / \log \log n \rceil$. Then

$$\begin{aligned} \mathbf{Pr}(A_{i,k}) &\leq \left(\frac{e}{k}\right)^k \frac{1}{1 - e/k} \\ &\leq 2\left(\frac{e}{3\log n/\log\log n}\right)^k \\ &\leq 2\left(e^{1 - \log 3 - \log\log n + \log\log\log n}\right)^k \\ &\leq 2\left(e^{-\log\log n + \log\log\log n}\right)^k \\ &\leq 2\left(e^{-3\log n + 3\frac{\log\log\log n}{\log\log n}\log n}\right) \\ &\leq 2\left(e^{-3\log n + 3\frac{\log\log\log n}{\log\log n}\log n}\right) \\ &\leq 2\left(e^{-2\log n}\right) \\ &= \frac{2}{n^2} \end{aligned}$$

for n sufficiently large that $(\log \log \log n) / \log \log n < 1/3$. It follows that

$$\mathbf{Pr}(\text{no bin contains more than } \lceil (3\log n)/\log\log n \rceil \text{ balls}) = 1 - \sum_{i=1}^{n} \mathbf{Pr}(A_{i,k})$$
$$\geq 1 - \frac{2}{n}$$

Theorem 6.1.1 Max Load

When n balls are thrown into n bins, the maximum number of balls in any bin is $O(\frac{\log n}{\log \log n})$ with high probability, i.e.,

$$E[max \ load] = \frac{\ln n}{\ln \ln n} (1 + o(1))$$
$$max \ load = \Theta(\frac{\ln n}{\ln \ln n}) \quad w.h.p.$$

It can be shown that this is a tight bound.

Coupon Collector's Problem Suppose I throw kn balls.

$$\mathbf{Pr}[\mathbf{bin \ 1 \ is \ empty}] \sim (\frac{1}{e})^k$$

If $k = c \ln n + d$, then

$$\mathbf{Pr}[\mathbf{bin 1 is empty}] \sim rac{1}{e^d n^c}$$

$$\mathbf{Pr}[\exists some \ bin \ empty] \le \frac{n}{n^c e^d} \le \frac{1}{n^{c-1}}$$

Therefore, w.h.p. $O(n \log n)$ balls suffice.

Claim:

$E[number of balls to see all bins] = n \cdot H_n$

Imagine a counter (starting at 0) that tells us how many boxes have at least one ball in it. Let X_1 denote the number of throws until the counter reaches 1 (so $X_1 = 1$). Let X_2 denote the number of throws from that point until the counter reaches 2. In general, let X_k denote the number of throws made from the time the counter hit k-1 up until the counter reaches k.

So, the total number of throws is $X_1 + ... + X_n$, and by linearity of expectation, what we are looking for is $E[X_1] + ... + E[X_n]$.

How to evaluate $E[X_k]$? Suppose the counter is currently at k-1. Each time we throw a ball, the probability it is something new is (n-(k-1))/n. So, another way to think about this question is as follows:

Coin flipping: we have a coin that has probability p of coming up heads (in our case, p = (n-(k-1))/n). What is the expected number of flips until we get a heads?

It turns out that the "intuitively obvious answer", 1/p, is correct. But why? Here is one way to see it: if the first flip is heads, then we are done; if not, then we are back where we started, except we've already paid for one flip. So the expected number of flips E satisfies: $E = p^*1 + (1-p)^*(1 + E)$. You can then solve for E = 1/p.

Putting this all together, let CC(n) be the expected number of throws until we have filled all the boxes. We then have:

$$CC(n) = E[X_1] + \dots + E[X_n]$$

= $n/n + n/(n-1) + n/(n-2) + \dots + n/1$
= $n(1/n + 1/(n-1) + \dots + 1/1)$
= nH_n

QED.

$$\mathbf{Pr}[x \ge n \ln n + cn \ or \ x \le n \ln n - cn] \backsim (e^{-e^{-c}} - e^{-e^{c}})$$

6.2 Hashing

FORMAL SETUP

- Keys come from some large universe M. (e.g., all < 50-character strings)
- Some set S in M of keys we actually care about (which may be static or dynamic).
- do inserts and lookups by having an array N of size |N|, and a HASH FUNCTION $h: M \to \{0, ..., |N| 1\}$. Given element x, store in N[h(x)].
- Will resolve collisions by having each entry in A be a linked list. Collision is when h(x) = h(y). There are other methods but this is cleanest called "separate chaining". To insert, just put at top of list. If h is good, then hopefully lists will be small.

UNIVERSAL HASHING

A hash family \mathcal{H} is 2-universal if for all $x \neq y$ in M,

$$\mathbf{Pr}_{h\in H}[h(x) = h(y)] \le \frac{1}{|N|}$$

Let $x, y \in M$.

$$C_{xy} = \begin{cases} 1 & \text{if } h(x) = h(y) \\ 0 & \text{otherwise} \end{cases}$$

$$E[C_{xy}] \le \frac{1}{|N|}$$

E[number of elts of S that collide with y] = $\sum_{x \neq y} C_{xy} \le \frac{|S|}{|N|}$

= E[amount of time when accessing y]

If $|N| \ge |S|$, then E[amount of time when accessing y] = o(1).

One way to construct a 2-universal hash family:

Here, let $M = \{0, ..., m-1\}$ and $N = \{0, ..., n-1\}$. Pick prime $p \ge m$ (or, think of just rounding m up to nearest prime). Define

$$h_{a,b}(x) = ((ax+b) \mod p) \mod n.$$

$$\mathcal{H} = \{h_{ab} | a, b \text{ in } GF(p) \text{ and } a \neq 0\}$$

It is easy to show that $|\mathcal{H}| = p(p-1)$.

Theorem 6.2.1 Lower Bound

 \mathcal{H} is a hash family $M \to N$, then $\exists x \neq y \in M$, s.t. $\mathbf{Pr}[h(x) = h(y)] \geq \frac{1}{|N|} - \frac{1}{|M|}$. Pf: via Yao's principle.

Strongly 2-univeral hash family see Anupam's notes

Perfect hash functions Definition: A hash function that maps each different key to a distinct integer. Usually all possible keys must be known beforehand. A hash table that uses a perfect hash has no collisions.

A family of hash functions $H = \{h : M \to N\}$ is said to be a perfect hash family if for each set $S \subseteq M$ of size $s \leq n$, there exists a hash function $h \in H$ that is perfect for S.

If |N| = |S|, every perfect hash family has size $2^{\Omega(|N|)}$.

2-level hashing [Fredman Komlos Szemerd]

Proposal: hash into table of size N. Will get some collisions. Then, for each bin, rehash it, squaring the size of the bin to get zero collisions.

To construct a 2-level hash function:

- 1. Pick $h \in H$, where H is a 2-universal hash family $M \to N$ and |N| = |S|.
- 2. If number of collisions > |N|, goto step 1
- 3. If N_i elements hashed to bin $i \leq N$, then pick $h_i : M \to N_i^2$. If any collisions go ostep 3.
- 4. Do step 3 for all bins.

$$\begin{aligned} \mathbf{Pr}[\mathbf{x}, \mathbf{y} \text{ collide}] &\leq \frac{1}{|N|} \\ E[\text{number of collisions}] &\leq \binom{|S|}{2} \frac{1}{|N|} \end{aligned}$$

1. In step 1 and 2, since |N| = |S|, let C denote number of collisions.

$$E[C] \le \binom{|S|}{2} \frac{1}{|S|} < \frac{|S|}{2}$$

According to Markov Inequality,

$$\Pr\left[C > 2 \cdot \frac{|S|}{2}\right] \le \frac{1}{2}$$

2. $C = \sum_{i} {N_i \choose 2} \le |N| = |S|$

3. If $H_i: M \to N_i^2$, set S is of size N_i .

$$E[C_i] = \leq \binom{N_i}{2} \cdot \frac{1}{N_i^2} \le \frac{1}{2}$$

Therefore, according to Markov Inequality,

$$\mathbf{Pr}[C_i \ge 1] \le \frac{1}{2}$$

Now let's study the space requirement of this scheme.

$$Space \leq |N| + \sum_i N_i^2 \leq 2|S|$$

In addition, to store the hash functions, we need to use O(|S|) more bits.

Unfortunately, this approach works for static dictionary only, but not dynamic dictionaries where we want to support insert/delete operations.